

GLOBAL RELIABILITY SENSITIVITY ANALYSIS OF MODELS WITH DEPENDENT INPUTS USING FAILURE SAMPLES Max Ehre, Iason Papaioannou, Daniel Straub

CONTRIBUTION: RSA WITH DEPENDENT INPUTS AS RA BYPRODUCT

Reliability sensitivity analysis (RSA) with dependent inputs is largely unexplored. We demonstrate how to efficiently compute the indices of [3] for reliability targets as byproduct of a single reliability analysis (RA).

Reliability analysis

The failure probability of a model with input a random vector(RV) $X \in \mathcal{X} \subseteq \mathbb{R}^d$, distribution $X \sim f_X$ and limit-state function (LSF) $g: \mathcal{X} \to \mathbb{R}$ is

$$\mathbb{P}(\mathbf{F}) = \int_{\mathcal{X}} \mathbf{I}[g(\boldsymbol{x}) \le 0] f_{\boldsymbol{X}}(\boldsymbol{x}) d\boldsymbol{x} = \mathbb{E}[\mathbf{I}[g(\boldsymbol{X}) \le 0]].$$
(1)

By convention, we choose the LSF to describe failure as $g(\boldsymbol{x}) \leq 0$, hence the integral over $f_{\boldsymbol{X}}$ censored on the failure domain $\mathbf{F} = \{ \boldsymbol{x} : g(\boldsymbol{x}) \leq 0 \}$ yields the failure probability. Given an isoprobabilistic transformation $T: \mathcal{X} \to \mathbb{R}^d$ with $U = T(\mathbf{X})$ and $\boldsymbol{U} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ (standard-normal RV), the reliability problems reads (φ is the standard-normal PDF)

$$\mathbb{P}(\mathbf{F}) = \int_{\mathbb{R}^d} \mathbf{I}[g(T^{-1}(\boldsymbol{u})) \le 0] \varphi(\boldsymbol{u}) d\boldsymbol{u}.$$
(2)

VARIANCE-BASED RSA

Variance-based sensitivity analysis (SA) of model output (Model $\mathcal{M} : \mathcal{X} \to \mathcal{Y}$ with $Y = \mathcal{M}(\mathbf{X})$) attributes fractions of the output variance $\mathbb{V}[Y]$ to the input variables $\{X_i\}_{i=1}^d$. After normalization with $\mathbb{V}[Y]$ we obtain first-order Sobol' indices as

$$S_{i} = \frac{\mathbb{V}[\mathbb{E}[Y|X_{i}]]}{\mathbb{V}[Y]}, \quad ST_{i} = 1 - \frac{\mathbb{V}[\mathbb{E}[Y|X_{\sim i}]]}{\mathbb{V}[Y]}. \quad (4)$$

 $\mathcal{M} = I[g(\mathbf{X}) \leq 0]$ yields reliability-oriented indices:

$$S_{i} = \frac{\mathbb{V}[\mathbb{E}[\mathrm{I}[g(\boldsymbol{X}) \leq 0] | X_{i}]]}{\mathbb{V}[\mathrm{I}[g(\boldsymbol{X}) \leq 0]]} \stackrel{(1)}{=} \frac{\mathbb{V}[\mathbb{P}(\mathrm{F}|X_{i})]}{\mathbb{V}[\mathrm{I}[g(\boldsymbol{X}) \leq 0]]} \quad (5)$$

for the Sobol' and likewise for the total Sobol' index. We can compute $\mathbb{P}(F|X_i)$ using Bayes' rule:

$$\mathbb{P}(\mathbf{F}|X_i) = \frac{f_{X_i}(x_i|\mathbf{F})\mathbb{P}(\mathbf{F})}{f_{X_i}(x_i)}.$$
 (6)

With sample-based RA methods (Monte Carlo, importance sampling, subset simulation), we get $\mathbb{P}(F)$ along with failure samples. These can be used to estimate $f_{X_i}(x_i|\mathbf{F})$ with kernel densities (KDE) [2].

For dependent inputs X, variance contributions stemming from \mathcal{M} and $f_{\mathbf{X}}$ are difficult to discern. A possible path forward is to transform the problem to an independent probability space.

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DEPENDENT INPUTS

- Variance contributions can be discerned by considering different transformations T from \mathcal{X} to standard-normal space.
- T is non-unique \rightarrow sensitivity metrics depend on the choice of T.
- Suppose, T has a hierarchical structure (e.g., the Nataf or the Rosenblatt transform):

$$T^{\{i\}}: \begin{bmatrix} U_i^{\{i\}} \\ U_{i+1}^{\{i\}} \\ \vdots \\ U_{i-1}^{\{i\}} \end{bmatrix} = \begin{bmatrix} h_1^{\{i\}}(X_i) \\ h_2^{\{i\}}(X_i, X_{i+1}) \\ \vdots \\ h_d^{\{i\}}(X_i, X_{i+1}, \dots, X_{i-1}) \end{bmatrix}$$

• [3] consider all d cyclic left shifts of the ordered set $\{X_1, \ldots, X_d\}$ (they use T^{-1} instead of T):



• Under $T^{\{i\}}, X_{i-1}$ $(X_0 \triangleq X_d)$ affects only $U_{i-1}^{\{i\}}$ and X_i affects all of $U^{\{i\}}$, hence one defines

$$\Rightarrow S_{i,\text{ind}} = \frac{\mathbb{V}[\mathbb{E}[Y|U_i^{\{i+1\}}]]}{\mathbb{V}[Y]}, \quad S_i = \frac{\mathbb{V}[\mathbb{E}[Y|U_i^{\{i\}}]]}{\mathbb{V}[Y]}.$$

- [3] use pick-freeze estimators in d repeated SA runs to obtain all independent and full Sobol'/total Sobol' indices.
- With n independent samples per SA run this requires a total of nd(d+1) model calls.





 $g(oldsymbol{X})$

We use failure samples from subset simulation (SUS) runs with $n = 10^4$ samples per level (3 levels) and compare with [1] & [3] each using $n = 10^6$ independent samples. Statistics are based on 100 repeated runs.



 $S_{i,in}^{i,jn}$

Our approach can be easily adpated to compute first-order independent and full indices of model output by using order statistics of the independent back-transformed output samplef or each transformation $T^{\{i\}}$ to computing $\mathbb{E}[Y|U_i^{\{1\}}]$.



EXAMPLE APPLICATION

We analyze the following LSF with $\mathbb{P}(F) = 6.2 \cdot 10^{-3}$.

$$= X_1^3 + 10X_2^2 + 0.1\sin(\pi X_2) + 10X_3^2 + 40\sin(\pi X_3) + 38, \text{ where } \boldsymbol{X} \sim \mathcal{N}\left(\begin{bmatrix}0\\0\\0\end{bmatrix}, \begin{bmatrix}1 & 0.5 & 0.3\\0.5 & 1 & 0.8\\0.3 & 0.8 & 1\end{bmatrix}\right).$$



OUTLOOK

[1]	S. Kucheren sitivity indi <i>Communicati</i>
[2]	Luyi Li, Ias sitivity estin 2019.
[3]	Thierry A. I methods for



$$\frac{\mathbb{P}[F|U_{i}^{\{i\}}]]}{[g(\mathbf{X}) \leq 0]]}.$$

$$(3)$$

$$\underbrace{\left\{F|U_{i}^{\{j\}}\right\}}_{i \in \{j-1,j\}} \underbrace{\mathbb{V}[.]}_{\{S_{i}\}_{i=1}^{d}} \underbrace{\{S_{i}\}_{i=1}^{d}}_{\{S_{i}\}_{i=1}^{d}} \underbrace{\{S_{i}\}_{i=1}^{d}}_{\{S_{i}\}_{i=1}^{d}} \underbrace{\{S_{i}\}_{i=1}^{d}}_{i \in \{j-1,j\}} \underbrace{\{u_{i}^{\{j\}}|F\}}_{i \in \{u_{i}^{\{j\}}|F\}} \underbrace{\{u_{i}^{\{j\}}|F\}}_{i \in \{u_{i}^{\{j\}}|F\}} \underbrace{\{u_{i}^{$$

Figure 1 (left): mean estimates (bar plot) $\pm \sigma$ (black brackets).

Table 1: Computational cost.

Method	LSF calls
Mara et al.	$12 \cdot 10^6$
Kucherenko	$3\cdot 10^6$
Failure samples	$0.03 \cdot 10^6$

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